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UDRC-EURASIP Summer School 2022

Heriot-Watt University

Problem



Problem

We observe $oldsymbol{X} \in \mathbb{R}^d$



Problem

We observe $oldsymbol{X} \in \mathbb{R}^d$

 $X\sim \mathbb{P}_0$



Problem

We observe $oldsymbol{X} \in \mathbb{R}^d$

 $X \sim \mathbb{P}_0$ or $X \sim \mathbb{P}_1$?



Problem We observe $oldsymbol{X} \in \mathbb{R}^d$ $oldsymbol{X} \sim \mathbb{P}_0$ or $oldsymbol{X} \sim \mathbb{P}_1$?

In classical *decision theory*, we *know* the distributions \mathbb{P}_0 and \mathbb{P}_1



Problem			
We observe	$oldsymbol{X}\in$	\mathbb{R}^{d}	
$oldsymbol{X}\sim\mathbb{P}_0$	or	$oldsymbol{X}\sim$	\mathbb{P}_1 ?

In classical *decision theory*, we *know* the distributions \mathbb{P}_0 and \mathbb{P}_1 In *machine learning*, we *have to estimate* \mathbb{P}_0 and \mathbb{P}_1 from data







$X \in \mathbb{R}$: number of **spam words** in a message



$X \in \mathbb{R}$: number of spam words in a message





$X \in \mathbb{R}$: number of spam words in a message

Null Hypothesis

 H_0 : message isn't spam





 $X \in \mathbb{R}$: number of spam words in a message

Null Hypothesis

 H_0 : message isn't spam

Alternative Hypothesis

 H_1 : message isn't spam



Consider a test for detecting:

if given email is spam

Consider a test for detecting:

if given email is spam

presence of aircraft in radar

Consider a test for detecting:

if given email is spam

if defendant is guilty

presence of aircraft in radar

Consider a test for detecting:

if given email is spampresence of aircraft in radarif defendant is guiltypresence of tumor in an image

Probability

Consider a test for detecting:

if given email is spam

if defendant is guilty

presence of aircraft in radar



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Consider a test for detecting:

if given email is spam pres

if defendant is guilty

presence of aircraft in radar presence of tumor in an image



Consider a test for detecting:

if given email is spam present

if defendant is guilty

presence of aircraft in radar presence of tumor in an image



Consider a test for detecting:

if given email is spam presence

if defendant is guilty

presence of aircraft in radar presence of tumor in an image



True hypothesis	Decide noise	Decide signal
noise		
signal		

True hypothesis	Decide noise	Decide signal
noise		
signal		



True hypothesis	Decide noise	Decide signal
noise	\checkmark	
signal		



True hypothesis	Decide noise	Decide signal
noise	\checkmark	
signal		\checkmark



True hypothesis	Decide noise	Decide signal
noise	\checkmark	false alarm
signal		\checkmark



True hypothesis	Decide noise	Decide signal
noise	\checkmark	false alarm
signal	missed detection	\checkmark


Larger effect size



Larger effect size





Decision Theory



Decision Theory

Where to place the decision boundary?





Decision Theory



True label

$$Y = \left\{ \begin{array}{rrr} 0 & , \mbox{ if } H_0 \mbox{ is true} \\ \\ 1 & , \mbox{ if } H_1 \mbox{ true} \end{array} \right.$$



True label

Decision function $f : \mathbb{R}^d \to \{0, 1\}$

 $Y = \begin{cases} 0 & , \text{ if } H_0 \text{ is true} \\ 1 & , \text{ if } H_1 \text{ true} \end{cases} \qquad f(\mathbf{X}) = \begin{cases} 0 & , \text{ if we } \underline{decide} \ H_0 \\ 1 & , \text{ if we } \underline{decide} \ H_1 \end{cases}$



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Loss function $\ell: \{0,1\} \times \{0,1\} \to \mathbb{R}$



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Loss function $\ell: \{0,1\} \times \{0,1\} \to \mathbb{R} \qquad \ell(f(\boldsymbol{X}),Y)$

True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
H_1 is true		

Decision Theory



True label

Decision function $f : \mathbb{R}^d \to \{0, 1\}$

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True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true	$\ell(0,0)$	
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True label

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True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	$\ell(0,0)$	$\ell(1,0)$
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H_0 is true	$\ell(0,0)$	$\ell(1,0)$
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Risk:

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Risk: $R[f] := \mathbb{E}_{XY} \Big[\ell \big(f(X), Y \big) \Big]$

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$$R[f] := \mathbb{E}_{XY} \Big[\ell \big(f(X), Y \big) \Big]$$

where $\mathbb{E}_{XY}[\cdot]$ is the expectation with respect to X and Y

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where $\mathbb{E}_{\boldsymbol{X}Y}[\cdot]$ is the expectation with respect to \boldsymbol{X} and Y

Optimal decision problem:

Given decision function $f: \mathbb{R}^d \to \{0,1\}$ and loss $\ell: \{0,1\}^2 \to \mathbb{R}$,

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... infinite-dimensional problem

Decision Theory

Risk:

 $\underset{f: \mathbb{R}^d \rightarrow \{0,1\}}{\text{minimize}} \ \mathbb{E}_{\boldsymbol{X}Y} \Big[\ell \big(f(\boldsymbol{X}), Y \big) \Big]$

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Recall that f(\mathbf{X}) and Y are binary
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Recall that $f(\boldsymbol{X})$ and Y are binary

Conditioning on X,

Recall that $f(\mathbf{X})$ and Y are binary Conditioning on \mathbf{X} ,

$$\mathbb{E}_{\boldsymbol{X}Y}\left[\ell(f(\boldsymbol{X},Y))\right] = \mathbb{E}_{\boldsymbol{X}}\left[\mathbb{E}_{Y}\left[\ell(f(\boldsymbol{X}),Y) \mid \boldsymbol{X}\right]\right]$$

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$$\mathbb{E}_{\boldsymbol{X}Y}\left[\ell(f(\boldsymbol{X},Y))\right] = \mathbb{E}_{\boldsymbol{X}}\left[\mathbb{E}_{Y}\left[\ell(f(\boldsymbol{X}),Y) \mid \boldsymbol{X}\right]\right]$$
$$= \int_{\mathbb{R}^{d}} \mathbb{E}_{Y}\left[\ell(f(\boldsymbol{X}),Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] f_{\boldsymbol{X}}(\boldsymbol{x}) \, \mathrm{d}\boldsymbol{x}$$

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$$\mathbb{E}_{\boldsymbol{X}\boldsymbol{Y}}\Big[\ell\big(f(\boldsymbol{X},\boldsymbol{Y})\big)\Big] = \mathbb{E}_{\boldsymbol{X}}\Big[\mathbb{E}_{\boldsymbol{Y}}\Big[\ell\big(f(\boldsymbol{X}),\,\boldsymbol{Y}\big)\,\Big|\,\boldsymbol{X}\Big]\Big]$$
$$= \int_{\mathbb{R}^d} \mathbb{E}_{\boldsymbol{Y}}\Big[\ell\big(f(\boldsymbol{X}),\,\boldsymbol{Y}\big)\,\Big|\,\boldsymbol{X} = \boldsymbol{x}\Big]f_{\boldsymbol{X}}(\boldsymbol{x})\,\mathrm{d}\boldsymbol{x}$$

If $f(\boldsymbol{x}) = 0$,

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If
$$f(\boldsymbol{x}) = 0$$
,

$$\mathbb{E}_{Y}\left[\ell(0, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] = \ell(0, 0) \mathbb{P}(Y = 0 \mid \boldsymbol{X} = \boldsymbol{x}) + \ell(0, 1) \mathbb{P}(Y = 1 \mid \boldsymbol{X} = \boldsymbol{x})$$

Recall that $f(\mathbf{X})$ and Y are binary Conditioning on \mathbf{X} ,

$$\mathbb{E}_{\boldsymbol{X}Y}\Big[\ell\big(f(\boldsymbol{X},Y)\big)\Big] = \mathbb{E}_{\boldsymbol{X}}\Big[\mathbb{E}_{Y}\Big[\ell\big(f(\boldsymbol{X}),Y\big) \,\Big|\, \boldsymbol{X}\Big]\Big]$$
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Recall that $f(\mathbf{X})$ and Y are binary Conditioning on \mathbf{X} ,

$$\mathbb{E}_{\boldsymbol{X}Y}\Big[\ell\big(f(\boldsymbol{X},Y)\big)\Big] = \mathbb{E}_{\boldsymbol{X}}\Big[\mathbb{E}_{Y}\Big[\ell\big(f(\boldsymbol{X}),Y\big) \mid \boldsymbol{X}\Big]\Big]$$
$$= \int_{\mathbb{R}^{d}} \mathbb{E}_{Y}\Big[\ell\big(f(\boldsymbol{X}),Y\big) \mid \boldsymbol{X} = \boldsymbol{x}\Big]f_{\boldsymbol{X}}(\boldsymbol{x}) \,\mathrm{d}\boldsymbol{x}$$

If $f(\mathbf{x}) = 0$, $\mathbb{E}_{Y}\left[\ell(0, Y) \mid \mathbf{X} = \mathbf{x}\right] = \ell(0, 0) \mathbb{P}(Y = 0 \mid \mathbf{X} = \mathbf{x}) + \ell(0, 1) \mathbb{P}(Y = 1 \mid \mathbf{X} = \mathbf{x})$ If $f(\mathbf{x}) = 1$, $\mathbb{E}_{Y}\left[\ell(1, Y) \mid \mathbf{X} = \mathbf{x}\right] = \ell(1, 0) \mathbb{P}(Y = 0 \mid \mathbf{X} = \mathbf{x}) + \ell(1, 1) \mathbb{P}(Y = 1 \mid \mathbf{X} = \mathbf{x})$

Decision Theory

Optimal decision

Optimal decision

$$f(\boldsymbol{x}) = 0$$
 if $\mathbb{E}_{Y}\left[\ell(0, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] < \mathbb{E}_{Y}\left[\ell(1, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right]$
$$f(\boldsymbol{x}) = 0 \quad \text{if} \quad \mathbb{E}_{Y}\left[\ell(0, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] < \mathbb{E}_{Y}\left[\ell(1, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right]$$
$$f(\boldsymbol{x}) = 1 \quad \text{if} \quad \mathbb{E}_{Y}\left[\ell(0, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] \ge \mathbb{E}_{Y}\left[\ell(1, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right]$$

$$\begin{split} f(\boldsymbol{x}) &= 0 \quad \text{ if } \quad \mathbb{E}_{Y} \bigg[\ell \big(0, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, < \, \mathbb{E}_{Y} \bigg[\ell \big(1, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \\ f(\boldsymbol{x}) &= 1 \quad \text{ if } \quad \mathbb{E}_{Y} \bigg[\ell \big(0, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, \ge \, \mathbb{E}_{Y} \bigg[\ell \big(1, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \end{split}$$

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$$f(\boldsymbol{x}) = \begin{cases} 1 & \text{if } \mathbb{P}(Y = 1 \mid \boldsymbol{X} = \boldsymbol{x}) \geq \frac{\ell(0,0) - \ell(1,0)}{\ell(1,1) - \ell(0,1)} \mathbb{P}(Y = 0 \mid \boldsymbol{X} = \boldsymbol{x}) \\ 0 & \text{otherwise} \end{cases}$$

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$$f(\boldsymbol{x}) = \begin{cases} 1 & \text{if } \mathbb{P}(\frac{\boldsymbol{Y} = 1}{\boldsymbol{Y} = 1} \mid \boldsymbol{X} = \boldsymbol{x}) \geq \frac{\ell(0,0) - \ell(1,0)}{\ell(1,1) - \ell(0,1)} \mathbb{P}(\boldsymbol{Y} = 0 \mid \boldsymbol{X} = \boldsymbol{x}) \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{split} f(\boldsymbol{x}) &= 0 \quad \text{ if } \quad \mathbb{E}_{Y} \bigg[\ell \big(0, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, < \, \mathbb{E}_{Y} \bigg[\ell \big(1, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \\ f(\boldsymbol{x}) &= 1 \quad \text{ if } \quad \mathbb{E}_{Y} \bigg[\ell \big(0, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, \ge \, \mathbb{E}_{Y} \bigg[\ell \big(1, \, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \end{split}$$

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$$\begin{aligned} f(\boldsymbol{x}) &= 0 \quad \text{if} \quad \mathbb{E}_{Y} \bigg[\ell \big(0, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, < \, \mathbb{E}_{Y} \bigg[\ell \big(1, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \\ f(\boldsymbol{x}) &= 1 \quad \text{if} \quad \mathbb{E}_{Y} \bigg[\ell \big(0, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \, \ge \, \mathbb{E}_{Y} \bigg[\ell \big(1, Y \big) \, \Big| \, \boldsymbol{X} = \boldsymbol{x} \bigg] \end{aligned}$$

$$f(\boldsymbol{x}) = \begin{cases} \begin{array}{c} H_1 \\ 1 & \text{if } \mathbb{P}(\underline{Y=1} \mid \boldsymbol{X} = \boldsymbol{x}) \geq \frac{\ell(0,0) - \ell(1,0)}{\ell(1,1) - \ell(0,1)} \mathbb{P}(\underline{Y=0} \mid \boldsymbol{X} = \boldsymbol{x}) \\ 0 & \text{otherwise} \end{cases} \\ \begin{array}{c} Bayes \ rule \end{cases}$$

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$$f(\boldsymbol{x}) = 0 \quad \text{if} \quad \mathbb{E}_{Y}\left[\ell(0, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right] < \mathbb{E}_{Y}\left[\ell(1, Y) \mid \boldsymbol{X} = \boldsymbol{x}\right]$$
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$$f(\boldsymbol{x}) = \begin{cases} \begin{array}{c} H_1 & H_0 \\ \text{if } \mathbb{P}(\underline{Y=1} \mid \boldsymbol{X} = \boldsymbol{x}) \geq \frac{\ell(0,0) - \ell(1,0)}{\ell(1,1) - \ell(0,1)} \mathbb{P}(\underline{Y=0} \mid \boldsymbol{X} = \boldsymbol{x}) \\ 0 & \text{otherwise} \\ & \underline{f_{\boldsymbol{X}\mid H_1}(\boldsymbol{x} \mid H_1) \mathbb{P}(H_1)} \\ \hline f_{\boldsymbol{X}}(\boldsymbol{x}) & \underline{f_{\boldsymbol{X}\mid H_0}(\boldsymbol{x} \mid H_0) \mathbb{P}(H_0)} \\ \end{array} \end{cases} \begin{vmatrix} H_0 & H_0 \\ H_0 \\ H_0 \\ \hline f_{\boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X} = \boldsymbol{X} \\ \hline f_{\boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X} \\ \hline f_{\boldsymbol{X}\mid \boldsymbol{X}\mid \boldsymbol{X}$$

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Rearranging, the optimal decision is

$$f(\boldsymbol{x}) = \begin{cases} 1 & \text{if } \mathbb{P}(\boldsymbol{Y}=\boldsymbol{1} \mid \boldsymbol{X}=\boldsymbol{x}) \geq \frac{\ell(0,0)-\ell(1,0)}{\ell(1,1)-\ell(0,1)} \mathbb{P}(\boldsymbol{Y}=\boldsymbol{0} \mid \boldsymbol{X}=\boldsymbol{x}) \\ 0 & \text{otherwise} \\ \frac{f_{\boldsymbol{X}\mid\boldsymbol{H}_{1}}(\boldsymbol{x}\mid\boldsymbol{H}_{1}) \mathbb{P}(\boldsymbol{H}_{1})}{f_{\boldsymbol{X}}(\boldsymbol{x})} \end{bmatrix} \begin{array}{l} \textbf{Bayes rule} \\ \frac{f_{\boldsymbol{X}\mid\boldsymbol{H}_{0}}(\boldsymbol{x}\mid\boldsymbol{H}_{0}) \mathbb{P}(\boldsymbol{H}_{0})}{f_{\boldsymbol{X}}(\boldsymbol{x})} \end{bmatrix}$$

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$$x - \frac{1}{2} = \log 0.76 \qquad \Longleftrightarrow \qquad x \simeq 0.23$$

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When class $Y \in \{0,1\}$ is viewed as a parameter of $\mathbb{P}_{\boldsymbol{X}Y}$ to estimate,

- Maximum a posteriori (MAP)
- Maximum likelihood (ML)

can be seen as likelihood ratio tests

 $\text{Consider} \ \ \ell(0,\,0)=\ell(1,\,1)=0 \ \ \text{and} \ \ \ell(1,\,0)=\ell(0,\,1)=1.$

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$. Then,

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$$f({m x})=1$$
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So, select
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 if $\mathbb{P}(Y = 1 \mid \boldsymbol{X} = \boldsymbol{x}) \geq \mathbb{P}(Y = 0 \mid \boldsymbol{X} = \boldsymbol{x})$

That is,
$$f(\boldsymbol{x}) = rg\max_i \frac{\mathbb{P}ig(Y=i \mid \boldsymbol{X}=\boldsymbol{x}ig)}{posterior}$$

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$$\mathbb{P}(Y = 1 \mid \boldsymbol{X} = \boldsymbol{x}) \geq \mathbb{P}(Y = 0 \mid \boldsymbol{X} = \boldsymbol{x})$$
$$\iff f_{\boldsymbol{X}\mid\boldsymbol{H}_{1}}(x \mid \boldsymbol{H}_{1}) \cdot \mathbb{P}(\boldsymbol{H}_{1}) \geq f_{\boldsymbol{X}\mid\boldsymbol{H}_{0}}(x \mid \boldsymbol{H}_{0}) \cdot \mathbb{P}(\boldsymbol{H}_{0})$$

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Recall that MAP rule minimizes probability of incorrect decision:
So, select
$$f(\boldsymbol{x}) = 1$$
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That is,
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This is a likelihood ratio test, because

$$\mathbb{P}(Y = 1 | \boldsymbol{X} = \boldsymbol{x}) \geq \mathbb{P}(Y = 0 | \boldsymbol{X} = \boldsymbol{x})$$

$$\iff f_{\boldsymbol{X}|H_1}(x | H_1) \cdot \mathbb{P}(H_1) \geq f_{\boldsymbol{X}|H_0}(x | H_0) \cdot \mathbb{P}(H_0)$$

$$\iff \mathcal{L}(\boldsymbol{x}) = \frac{f_{\boldsymbol{X}|H_1}(x | H_1)}{f_{\boldsymbol{X}|H_0}(x | H_0)} \geq \frac{\mathbb{P}(H_0)}{\mathbb{P}(H_1)} = \eta$$

Recall that MAP rule minimizes probability of incorrect decision: $\mathbb{P}(\text{error}) = \mathbb{P}(f(\mathbf{X}) = 1, H_0) + \mathbb{P}(f(\mathbf{X}) = 0, H_1)$

Decision Theory

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$

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$$f(\boldsymbol{x}) = \underset{i}{\operatorname{arg\,max}} \mathbb{P}(Y = i \mid \boldsymbol{X} = \boldsymbol{x})$$

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$ And $\mathbb{P}(H_0) = \mathbb{P}(H_1) = \frac{1}{2}$

$$\begin{split} f(\boldsymbol{x}) &= \arg \max_{i} \quad \mathbb{P} \big(Y = i \, \big| \, \boldsymbol{X} = \boldsymbol{x} \big) \\ &= \arg \max_{i} \quad \frac{f_{\boldsymbol{X}|Y}(\boldsymbol{x} \,| \, Y = i) \cdot \mathbb{P}(Y = i)}{f_{\boldsymbol{X}}(\boldsymbol{x})} \quad (\text{Bayes rule}) \end{split}$$

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$ And $\mathbb{P}(H_0) = \mathbb{P}(H_1) = \frac{1}{2}$

$$\begin{split} f(\boldsymbol{x}) &= \arg \max_{i} \quad \mathbb{P}\left(Y = i \mid \boldsymbol{X} = \boldsymbol{x}\right) \\ &= \arg \max_{i} \quad \frac{f_{\boldsymbol{X}|Y}(\boldsymbol{x} \mid Y = i) \cdot \mathbb{P}(Y = i)}{f_{\boldsymbol{X}}(\boldsymbol{x})} \qquad \qquad \left(\text{Bayes rule} \right) \\ &= \arg \max_{i} \quad f_{\boldsymbol{X}|Y}\left(\boldsymbol{x} \mid Y = i\right) \qquad \qquad \left(\mathbb{P}\left(Y = i\right) = \frac{1}{2} \right) \end{split}$$

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$ And $\mathbb{P}(H_0) = \mathbb{P}(H_1) = \frac{1}{2}$

$$\begin{split} f(\boldsymbol{x}) &= \arg \max_{i} \quad \mathbb{P}\left(Y = i \mid \boldsymbol{X} = \boldsymbol{x}\right) \\ &= \arg \max_{i} \quad \frac{f_{\boldsymbol{X}|Y}(\boldsymbol{x} \mid Y = i) \cdot \mathbb{P}(Y = i)}{f_{\boldsymbol{X}}(\boldsymbol{x})} \qquad \qquad \left(\text{Bayes rule} \right) \\ &= \arg \max_{i} \quad f_{\boldsymbol{X}|Y}\left(\boldsymbol{x} \mid Y = i\right) \qquad \qquad \left(\mathbb{P}\left(Y = i\right) = \frac{1}{2} \right) \\ &= \max_{i} \min \text{ maximum likelihood} \end{split}$$

Consider $\ell(0, 0) = \ell(1, 1) = 0$ and $\ell(1, 0) = \ell(0, 1) = 1$ And $\mathbb{P}(H_0) = \mathbb{P}(H_1) = \frac{1}{2}$

The optimal decision (MAP) is

$$\begin{split} f(\boldsymbol{x}) &= \arg \max_{i} \quad \mathbb{P}\left(Y = i \mid \boldsymbol{X} = \boldsymbol{x}\right) \\ &= \arg \max_{i} \quad \frac{f_{\boldsymbol{X}|Y}(\boldsymbol{x} \mid Y = i) \cdot \mathbb{P}(Y = i)}{f_{\boldsymbol{X}}(\boldsymbol{x})} \qquad \qquad \left(\text{Bayes rule} \right) \\ &= \arg \max_{i} \quad f_{\boldsymbol{X}|Y}\left(\boldsymbol{x} \mid Y = i\right) \qquad \qquad \left(\mathbb{P}\left(Y = i\right) = \frac{1}{2} \right) \\ &= \max \text{maximum likelihood} \end{split}$$

This corresponds to a likelihood ratio test with $\eta=1$

Decision Theory

	Table	of probabilities
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
H_1 is true		

True Positive Rate (TPR)

	Table	of probabilities
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
H_1 is true		

True Positive Rate (TPR)

power, sensitivity, recall

	Table	of probabilities
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
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True Positive Rate (TPR)

power, sensitivity, recall

	Table	of probabilities
True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
H_1 is true		TPR

True Positive Rate (TPR)

$$\mathsf{TPR} = \mathbb{P}\big(f(\boldsymbol{X}) = 1 \,\big|\, \boldsymbol{H}_1\big)$$

power, sensitivity, recall

True hypothesis	Table of probabilities	
	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
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True Positive Rate (TPR)

$$\mathsf{TPR} = \mathbb{P}(f(\boldsymbol{X}) = 1 \mid \boldsymbol{H}_1)$$

power, sensitivity, recall

False Positive Rate (FPR)

type I error, false alarm

	Table	of probabilities
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		
H_1 is true		TPR

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$$\mathsf{TPR} = \mathbb{P}(f(\boldsymbol{X}) = 1 \mid \boldsymbol{H}_1)$$

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False Positive Rate (FPR)

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True hypothesis	Table of probabilities	
	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true		FPR
H_1 is true		TPR

True Positive Rate (TPR)

power, sensitivity, recall

False Positive Rate (FPR)

type I error, false alarm

$$\mathsf{TPR} = \mathbb{P}(f(\mathbf{X}) = 1 \mid \mathbf{H}_1)$$

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True hypothesis	Table of probabilities		
	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$	
H_0 is true		FPR	
H_1 is true		TPR	

True Positive Rate (TPR)

power, sensitivity, recall

False Positive Rate (FPR)

type I error, false alarm

True Negative Rate (TNR) *specificity*

$$\mathsf{TPR} = \mathbb{P}\big(f(\boldsymbol{X}) = 1 \,\big|\, \boldsymbol{H}_1\big)$$

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True hypothesis	Table of probabilities		
	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$	
H_0 is true		FPR	
H_1 is true		TPR	

True Positive Rate (TPR)

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	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$	
H_0 is true	TNR	FPR	
H_1 is true		TPR	

True Positive Rate (TPR)

power, sensitivity, recall

False Positive Rate (FPR) type I error, false alarm

True Negative Rate (TNR) *specificity*

$$\mathsf{TPR} = \mathbb{P}(f(\mathbf{X}) = 1 \mid \mathbf{H}_1)$$

$$\mathsf{FPR} = \mathbb{P}\big(f(\boldsymbol{X}) = 1 \,\big|\, H_0\big)$$

$$\mathsf{TNR} = \mathbb{P}\big(f(\boldsymbol{X}) = 0 \,\big|\, H_0\big)$$

	Table of probabilities	
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	FPR
H_1 is true		TPR

True Positive Rate (TPR)

power, sensitivity, recall

False Positive Rate (FPR) *type I error, false alarm*

True Negative Rate (TNR) *specificity*

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$TPR = \mathbb{P}(f(\mathbf{X}))$	= 1	$ H_1)$
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$$\mathsf{FPR} = \mathbb{P}\big(f(\boldsymbol{X}) = 1 \,\big|\, H_0\big)$$

$$\mathsf{TNR} = \mathbb{P}\big(f(\boldsymbol{X}) = 0 \,\big|\, H_0\big)$$

	Table of probabilities	
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	FPR
H_1 is true		TPR

True Positive Rate (TPR)

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	Table of probabilities	
True hypothesis	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	FPR
H_1 is true	FNR	TPR

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$$\mathsf{TPR} = \mathbb{P}\big(f(\boldsymbol{X}) = 1 \,\big|\, \boldsymbol{H}_1\big)$$

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True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	FPR
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True Positive Rate (TPR) *power, sensitivity, recall*

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	Table of probabilities	
True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	$^{lpha}_{ m FPR}$
H_1 is true	FNR	TPR

True Positive Rate (TPR)

power, sensitivity, recall

False Positive Rate (FPR) *type I error, false alarm*

True Negative Rate (TNR) *specificity*

False Negative Rate (FNR)

$$\mathsf{TPR} = \mathbb{P}(f(\boldsymbol{X}) = 1 \mid \boldsymbol{H}_1)$$

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	Table of probabilities	
True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	$^{lpha}_{ m FPR}$
H_1 is true		TPR

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<u>Alternatives</u>

Precision: $\mathbb{P}(H_1 | f(X) = 1)$

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Precision: $\mathbb{P}(H_1 \mid f(X) = 1) = \frac{\mathsf{TPR} \cdot \mathbb{P}(H_1)}{\mathsf{TPR} \cdot \mathbb{P}(H_1) + \mathsf{FPR} \cdot \mathbb{P}(H_0)}$

 F_1 -score: harmonic mean between precision and recall (TPR):

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Precision:
$$\mathbb{P}(H_1 \mid f(X) = 1) = \frac{\mathsf{TPR} \cdot \mathbb{P}(H_1)}{\mathsf{TPR} \cdot \mathbb{P}(H_1) + \mathsf{FPR} \cdot \mathbb{P}(H_0)}$$

 F_1 -score: harmonic mean between precision and recall (TPR):

$$F_1 = \frac{2}{\frac{1}{\mathbb{P}(H_1 \mid f(X) = 1)} + \frac{1}{\mathsf{TPR}}}$$

	Table of probabilities	
True hypothesis	$f(\pmb{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	$^{\alpha}$ FPR
H_1 is true		TPR
True hypothesis	Table of probabilities	
-----------------	-------------------------	-------------------------
	$f(\boldsymbol{X}) = 0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	lpha FPR
H_1 is true	β FNR	TPR

 α and β are in *conflict*:

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True hypothesis	$f(\boldsymbol{X})=0$	$f(\boldsymbol{X}) = 1$
H_0 is true	TNR	lpha FPR
H_1 is true	FNR B	TPR

 α and β are in conflict:

•
$$\alpha \downarrow \implies \beta \uparrow$$
 : and vice-versa

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- $\alpha \downarrow \implies \beta \uparrow$: and vice-versa
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H_0 is true	TNR	lpha FPR
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 α and β are in *conflict*:

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It turns out that likelihood ratio tests are Pareto optimal

Theorem

Let $f_{\mathsf{LRT}}(\boldsymbol{x})$ be a likelihood ratio decision rule

Theorem

Let $f_{\mathsf{LRT}}({m{x}})$ be a likelihood ratio decision rule with FPR and FNR

$$\alpha_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 1 \,\big|\, H_0\Big) \qquad \beta_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 0 \,\big|\, \boldsymbol{H}_1\Big)$$

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Let f(x) be another (deterministic or probabilistic) decision rule with

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Let $f(\boldsymbol{x})$ be another (deterministic or probabilistic) decision rule with

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Theorem

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Let $f(\boldsymbol{x})$ be another (deterministic or probabilistic) decision rule with

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Then,

$$\alpha \leq \alpha_{\mathsf{LRT}} \qquad \Longrightarrow \qquad \beta \geq \beta_{\mathsf{LRT}}$$

Theorem

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$$\alpha_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 1 \,\big|\, H_0\Big) \qquad \beta_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 0 \,\big|\, \boldsymbol{H}_1\Big)$$

Let $f(\boldsymbol{x})$ be another (deterministic or probabilistic) decision rule with

$$\alpha = \mathbb{P}\Big(f(\boldsymbol{x}) = 1 \mid H_0\Big) \qquad \beta = \mathbb{P}\Big(f(\boldsymbol{x}) = 0 \mid \boldsymbol{H}_1\Big)$$

Then,

$\alpha \leq \alpha_{\text{LRT}}$	\Rightarrow	$\beta\geq\beta_{\rm LRT}$
$\beta \leq \beta_{\rm LRT}$	\implies	$\alpha \geq \alpha_{\rm LRT}$

Theorem

Let $f_{\mathsf{LRT}}(\pmb{x})$ be a likelihood ratio decision rule with FPR and FNR

$$\alpha_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 1 \,\big|\, H_0\Big) \qquad \beta_{\mathsf{LRT}} = \mathbb{P}\Big(f_{\mathsf{LRT}}(\boldsymbol{x}) = 0 \,\big|\, H_1\Big)$$

Let $f(\boldsymbol{x})$ be another (deterministic or probabilistic) decision rule with

$$\alpha = \mathbb{P}\Big(f(\boldsymbol{x}) = 1 \mid H_0\Big) \qquad \beta = \mathbb{P}\Big(f(\boldsymbol{x}) = 0 \mid \boldsymbol{H}_1\Big)$$

Then,

$\alpha\leq\alpha_{\rm LRT}$	\implies	$\beta\geq\beta_{\rm LRT}$
$\beta\leq\beta_{\rm LRT}$	\implies	$\alpha \geq \alpha_{\rm LRT}$

And the same relations hold with strict inequalities (<,>)

Decision Theory

$$\begin{split} & \mathbb{P}\Big(\mathsf{error}_{\mathsf{MAP}}\Big) \\ & = \mathbb{P}\Big(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1, \, H_0\Big) + \mathbb{P}\Big(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1, \, H_1\Big) \end{split}$$

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$$\mathbb{P}\left(\mathsf{error}_{\mathsf{MAP}}\right)$$

$$= \mathbb{P}\left(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1, H_0\right) + \mathbb{P}\left(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1, H_1\right)$$

$$= \underbrace{\mathbb{P}\left(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1 \mid H_0\right)}_{\alpha_{\mathsf{MAP}}} \mathbb{P}(H_0) + \mathbb{P}\left(f_{\mathsf{MAP}}(\boldsymbol{X}) = 1 \mid H_1\right) \mathbb{P}(H_1)$$

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This implies $f_{\text{LRT}}(\boldsymbol{x}) = f_{\text{MAP}}(\boldsymbol{x})$

Decision Theory

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Receiver Operating Characteristic (ROC)

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Decision Theory

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$$f_{X|H_0}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x}{2\sigma}}$$





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 $\mathsf{TPR}(\eta)$

$$x \ge \frac{c}{2} + \frac{\sigma^2}{c} \log \eta := \gamma$$

$$\mathsf{TPR}(\eta) = \mathbb{P}\Big(X \ge \gamma \, \big| \, \underline{H}_1\Big)$$

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$$\mathsf{TPR}(\eta) = \mathbb{P}\Big(X \ge \gamma \, \big| \, H_1\Big) = \int_{\gamma}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-c)^2}{2\sigma^2}} \, \mathrm{d} \, x$$

$$x \ge \frac{c}{2} + \frac{\sigma^2}{c} \log \eta := \gamma$$

$$\begin{aligned} \mathsf{TPR}(\eta) &= \mathbb{P}\Big(X \ge \gamma \, \big| \, H_1\Big) = \int_{\gamma}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-c)^2}{2\sigma^2}} \, \mathrm{d}\, x \\ &= \int_{\frac{\gamma-c}{\sigma}}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} \, \, \mathrm{d}z \end{aligned}$$

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Similarly,

$$\mathsf{FPR}(\eta) = \mathbb{P}\left(X \ge \gamma \mid H_0\right) = \cdots = Q\left(\frac{\log \eta}{\mathsf{SNR}} + \frac{\mathsf{SNR}}{2}\right)$$

ROC curve for different values of SNR

ROC curve for different values of SNR

 $\mathsf{TPR}(\eta)$ 1.0 -0.8 0.6 0.4 0.2 -0.0 + 0.4 0.0 0.2 0.6 0.8 1.0 $\mathsf{FPR}(\eta)$





Decision Theory





Decision Theory





Decision Theory

ROC curve for different values of SNR



Decision Theory
Example

ROC curve for different values of SNR



Decision Theory

Property 1: (0,0) and (1,1) are in the ROC curve

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Property 2: TPR \geq FPR

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Proof:

• When $\eta \to +\infty$,

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for any \boldsymbol{x}

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• Similarly, $(FPR(-\infty), TPR(-\infty)) = (1,1)$

Consider two decision rules (omitting dependence on η):

 $\left(\mathsf{FPR}^{(1)} \,,\, \mathsf{TPR}^{(1)} \right) \ , \ \left(\mathsf{FPR}^{(2)} \,,\, \mathsf{TPR}^{(2)} \right)$

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$$\left(\mathsf{FPR}\,,\,\mathsf{TPR}\right) = \left(p\,\mathsf{FPR}^{(1)} + (1-p)\mathsf{FPR}^{(2)}\,,\,\,p\,\mathsf{TPR}^{(1)} + (1-p)\mathsf{TPR}^{(2)}\right)$$

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Lab-based tests (ELISA, CLIA) and rapid tests (lateral flow)

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Decision Theory

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Decision Theory

We studied a (binary) decision problem:

$$\underset{f:\mathbb{R}^d \rightarrow \{0,1\}}{\text{minimize}} \ \mathbb{E}_{\boldsymbol{X}Y}\Big[\ell\big(f(\boldsymbol{X}),Y\big)\Big]$$

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Decision Theory

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Optimal decision problems (binary case)



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Likelihood ratio tests (LRT)



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MAP and ML as particular cases



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ROC curves and properties

Probability decision boundary

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References

References



M. Hardt, B. Recht

Patterns, Predictions, and Actions

Princeton University Press, Oct, 2022

References



M. Hardt, B. Recht Patterns, Predictions, and Actions Princeton University Press, Oct, 2022



D. P. Bertsekas, J. N. Tsitsiklis Introduction to Probability

Athena Scientific, 2nd edition, 2008

T. Fauwcett, "An introduction to ROC analysis," Pattern Recognition Letters,

Vol. 27, pp. 861-874, 2006.

H. L. Van Trees, **Detection, Estimation, and Modulation: Part I**, John Wiley & Sons, 2001.